

Final Exam

1. (34 points) State whether each of the following is true or false. [No Explanation Necessary]
 - (a) The larger the sample size, the smaller the variance of the sample standard deviation.
Answer: T
 - (b) The central limit theorem is not applied to discrete random variables.
Answer: F
 - (c) The sample standard deviation is an example of an estimator.
Answer: T
 - (d) The p-value of two-sided test is not equal to the p-value of one-sided test.
Answer: T
 - (e) If the p-value of a two sided test for the mean of a population is .005, then the null hypothesis will be rejected at a level of significance of .01.
Answer: T
 - (f) If the null hypothesis is not rejected based on sample evidence, the researcher has proven beyond a doubt that the null hypothesis is true.
Answer: F
 - (g) If a null hypothesis is rejected against an alternative at the 5% level, then using the same data, it must be rejected against that alternative at the 10% level.
Answer: T
 - (h) As the sample size increases to infinity, the variance of the sample mean approaches zero.
Answer: T

2. Let $X_1, X_2, X_3,$ and X_4 be a random sample of observations from a population with mean μ and variance σ^2 . Consider the following two point estimators of μ : $\hat{\theta}_1 = 0.25X_1 + 0.25X_2 + 0.40X_3 + 0.10X_4$ and $\hat{\theta}_2 = 0.20X_1 + 0.30X_2 + 0.30X_3 + 0.20X_4$.
 - (a) (4 points) Which of the following is true? [No Explanation Necessary]
 - (1) $\hat{\theta}_1$ is biased, but $\hat{\theta}_2$ is unbiased estimator of μ . (2) $\hat{\theta}_1$ is unbiased, but $\hat{\theta}_2$ is biased estimator of μ . (3) Both $\hat{\theta}_1$ and $\hat{\theta}_2$ are unbiased estimators of μ . (4) Both $\hat{\theta}_1$ and $\hat{\theta}_2$ are biased estimators

of μ .

ANSWER: (3)

(b) (4 points) Which of the following is true? [No Explanation Necessary]

(1) $Var(\hat{\theta}_1) = Var(\hat{\theta}_2)$. (2) $Var(\hat{\theta}_1) > Var(\hat{\theta}_2)$. (3) $Var(\hat{\theta}_1) < Var(\hat{\theta}_2)$. (4) Cannot tell the relationship between $Var(\hat{\theta}_1)$ and $Var(\hat{\theta}_2)$.

ANSWER: $Var(\hat{\theta}_1) = 0.295 > Var(\hat{\theta}_2) = 0.26$ so that (2) is the answer.

3. (6 points) Assume that the yield per acre for a particular variety of soybeans is $N(\mu, \sigma^2)$. For a random sample of $n = 4$ plots, the yields per acre were 37.4, 48.8, 46.9, and 55.0. Find a 90 percent confidence interval for μ .

ANSWER: $\bar{X} = (1/4)(37.4 + 48.8 + 46.9 + 55) = 47.025$ and $s^2 = (1/3)((37.4 - 47.025)^2 + (48.8 - 47.025)^2 + (46.9 - 47.025)^2 + (55 - 47.025)^2) = 53.13$ and $s = 7.29$. From the table for the Student's t distribution when d.f. = 3, we have $t_{0.05} = 2.353$. Therefore, a 90 percent confidence interval is given by $\bar{X} \pm 2.353 \times (s/\sqrt{4}) = 47.03 \pm 8.58$ or (38.45, 48.61).

Grading Note: if s^2 is computed by dividing by 4 rather than 3, give zero point. If $z_{0.05} = 1.64$ or 1.65 (or any other numbers such as $t_{0.1} = 1.533$) is used in place of $t_{0.05} = 2.353$, then give zero point. If students cannot find $t_{0.05}$ from t-table, give zero point. Any other clear computational mistake (perhaps, typing incorrect numbers in computing \bar{X} or s), give 3 points.

4. If a newborn baby has a birth weigh that is less than 2500 grams, we say that the baby has a low birth weight. The proportion of babies with a low birth weight is an indicator of lack of nutrition for the mothers. In Canada, approximately 7 percent of babies have a low birth weight. Let p equal the proportion of babies born in the Sudan who weight less than 2500 grams. In a random sample of $n = 209$ babies in the Sudan, there are 23 babies whose weight is less than 2500 grams.

(a) (6 points) Test the null hypothesis $H_0 : p \leq 0.07$ against the alternative hypothesis $H_1 : p > 0.07$ at a significance level of $\alpha = 0.05$.

ANSWER: $\hat{p} = 0.11$ and $\sqrt{\hat{V}ar(\hat{p})} = \sqrt{\hat{p}(1 - \hat{p})/209} = \sqrt{0.11 \times 0.89/209} = 0.0216$. Under H_0 , \hat{p} is approximately distributed as normal distribution with mean $p_0 = 0.07$ and standard deviation $\sqrt{p_0(1 - p_0)/n} = \sqrt{0.07 \times (1 - 0.07)/209} = 0.0176$. The null hypothesis is rejected if the test statistic $\frac{\hat{p} - p_0}{\sqrt{p_0(1 - p_0)/n}}$ is greater than $z_{0.05} = 1.64$ or 1.65. Here, $\frac{\hat{p} - p_0}{\sqrt{p_0(1 - p_0)/n}} = \frac{0.11 - 0.07}{0.0176} = 2.273$. Therefore, H_0 is rejected.

Grading Note: If the student uses $\sqrt{\hat{p}(1 - \hat{p})/209} = \sqrt{0.11 \times 0.89/209} = 0.0216$ in place

of $\sqrt{p_0(1-p_0)/n} = \sqrt{0.07 \times (1-0.07)/209} = 0.0176$ to compute the standard deviation for \hat{p} , this is OK and give full points because it is not wrong. Any other answer get zero points unless it is clear that a student makes computational mistake (say, typing wrong numbers into calculator), in which case student will get 3 points.

- (b) (6 points) Find the p-value for this test.

ANSWER: In Standard Normal distribution table, $\Pr(z > 2.27) = 0.5 - 0.4884 = 0.0116$ and $\Pr(z > 2.28) = 0.5 - 0.4887 = 0.0113$. Therefore, the p-value is between 0.0113 and 0.0116.

Grading Note: both 0.0113 and 0.0116 are acceptable for the answer. Any other numbers will get zero point.

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5. Let X , Y , and Z are three Bernoulli random variables that are independent to each other with the probability of success p_x , p_y , and p_z , respectively. We are interested in testing the null hypothesis $H_0 : p_x + p_y \geq p_z$ against the alternative hypothesis $H_1 : p_x + p_y < p_z$. Suppose that we have a random sample of X , Y , and Z with sample size $n = 100$: $\{X_1, \dots, X_n\}$, $\{Y_1, \dots, Y_n\}$, and $\{Z_1, \dots, Z_n\}$. The sample averages (i.e., sample proportions) of X , Y , and Z are given by $\hat{p}_x = (1/n) \sum_{i=1}^n X_i = 0.1$, $\hat{p}_y = (1/n) \sum_{i=1}^n Y_i = 0.3$, and $\hat{p}_z = (1/n) \sum_{i=1}^n Z_i = 0.5$. Denote the true value of p_x , p_y , and p_z by p_x^0 , p_y^0 , and p_z^0 ,

- (a) (6 points) What is the estimated variance of $\hat{p}_x + \hat{p}_y - \hat{p}_z$?

ANSWER: $Var(\hat{p}_x + \hat{p}_y - \hat{p}_z) = Var(\hat{p}_x) + Var(\hat{p}_y) + Var(\hat{p}_z) = p_x^0(1-p_x^0)/n + p_y^0(1-p_y^0)/n + p_z^0(1-p_z^0)/n$, where the first equality follows from $Var(\hat{p}_x + \hat{p}_y - \hat{p}_z) = Var(\hat{p}_x) + Var(\hat{p}_y) + Var(\hat{p}_z) + 2Cov(\hat{p}_x, \hat{p}_y) - 2Cov(\hat{p}_x, \hat{p}_z) - 2Cov(\hat{p}_y, \hat{p}_z)$ (prove this yourself), where $Cov(\hat{p}_x, \hat{p}_y) = Cov(\hat{p}_x, \hat{p}_z) = Cov(\hat{p}_y, \hat{p}_z) = 0$ by random sampling assumption and the assumption that X , Y , and Z are independent to each other. Using the estimator of \hat{p}_x , \hat{p}_y , and \hat{p}_z in place of p_x^0 , p_y^0 , and p_z^0 , we have $\hat{V}ar(\hat{p}_x + \hat{p}_y - \hat{p}_z) = (1/100) * (0.1 * 0.9 + 0.3 * 0.7 + 0.5 * 0.5) = 0.0055$. If the answer is given by $Var(\hat{p}_x + \hat{p}_y - \hat{p}_z) = Var(\hat{p}_x) + Var(\hat{p}_y) + Var(\hat{p}_z) = p_x^0(1-p_x^0)/n + p_y^0(1-p_y^0)/n + p_z^0(1-p_z^0)/n$ rather than 0.0055, give a full mark of 6 points.

Grading Note: No point unless students get 0.0055 unless it is clear that student understand that $Var(\hat{p}_x + \hat{p}_y - \hat{p}_z) = Var(\hat{p}_x) + Var(\hat{p}_y) + Var(\hat{p}_z) = p_x^0(1-p_x^0)/n + p_y^0(1-p_y^0)/n + p_z^0(1-p_z^0)/n$ and just make mistake in computing 0.0055.

- (b) (6 points) Test the null hypothesis $H_0 : p_x + p_y \geq p_z$ against the alternative hypothesis $H_1 : p_x + p_y < p_z$ at the significance level $\alpha = 0.1$. [Hint: you can use sample proportions \hat{p}_x , \hat{p}_y , and \hat{p}_z in place of p_x^0 , p_y^0 , and p_z^0 to compute the standard deviation for $\hat{p}_x + \hat{p}_y - \hat{p}_z$ when you construct a test statistic.]

ANSWER: Under the null hypothesis $H_0 : p_x + p_y - p_z = 0$, the test statistic $z = \frac{\hat{p}_x + \hat{p}_y - \hat{p}_z - 0}{\sqrt{\text{Var}(\hat{p}_x + \hat{p}_y - \hat{p}_z)}}$ will be approximately distributed as $N(0, 1)$. We reject the null hypothesis if the realized value of $z = \frac{\hat{p}_x + \hat{p}_y - \hat{p}_z - 0}{\sqrt{\text{Var}(\hat{p}_x + \hat{p}_y - \hat{p}_z)}}$ is smaller than $z_{0.9} = -1.28$. Using 0.0055 as an estimate for $\text{Var}(\hat{p}_x + \hat{p}_y - \hat{p}_z)$, we have $\frac{\hat{p}_x + \hat{p}_y - \hat{p}_z - 0}{\sqrt{\text{Var}(\hat{p}_x + \hat{p}_y - \hat{p}_z)}} = \frac{-0.1}{\sqrt{0.0055}} = -1.35$. Because -1.35 is smaller than -1.28 , we reject H_0 . **Grading Note:** If student construct the test statistic $z = \frac{\hat{p}_x + \hat{p}_y - \hat{p}_z - 0}{\sqrt{\text{Var}(\hat{p}_x + \hat{p}_y - \hat{p}_z)}}$, give 3 points. A full mark requires (i) computing critical value -1.28 and computing the test statistic's value -1.35 and correctly state that H_0 is rejected.

6. Let $\{X_1, X_2, \dots, X_n\}$ be $n = 9$ observations, each of which is randomly drawn from normal distribution with mean μ and variance σ^2 . The value of μ is not known while σ^2 is known and equal to 25. We are interested in testing the null hypothesis $H_0 : \mu = 10$ against the alternative hypothesis $H_1 : \mu < 10$. Consider hypothesis testing based on the following two different test statistics: (i) sample mean $\bar{X} = (1/n) \sum_{i=1}^n X_i$ and (ii) weighted average $\hat{X} = 0.6X_1 + 0.05 \times \sum_{i=2}^9 X_i$. Suppose that the realized values of \bar{X} and \hat{X} are given by $\bar{X} = 7.5$ and $\hat{X} = 6.5$, respectively.

- (a) (6 points) Test the null hypothesis $H_0 : \mu = 10$ against $H_1 : \mu < 10$ using the test statistic \hat{X} at the significance level $\alpha = 0.1$. Similarly, test the null hypothesis using the test statistic \bar{X} .

Answer: $\sqrt{\text{Var}(\bar{X})} = \sqrt{25/9} \approx 1.667$ and $\sqrt{\text{Var}(\hat{X})} = \sqrt{(0.6)^2 + 8 \times (0.05)^2} \times 5 = \sqrt{0.38} \times 5 \approx 3.082$. The test using \bar{X} will reject H_0 if \bar{X} is less than $10 - 1.28 \times 1.667 = 7.866$ so that, given $\bar{X} = 7.5$, and the null hypothesis is rejected using \bar{X} at $\alpha = 0.1$ because $7.5 < 7.866$. The test using \hat{X} will reject H_1 if \hat{X} is less than $10 - 1.28 \times 3.082 = 6.055$, and the null hypothesis is not rejected using \hat{X} at $\alpha = 0.1$ because $6.5 > 6.055$.

Grading Note: 3 points for each of sub-questions.

- (b) (6 points) Compute (i) the power of test using using the test statistic \bar{X} when the true value of μ is equal to 8 and (ii) the power of test using the test statistic \hat{X} when the true value of μ is equal to 8. Based on the power comparison, which test statistics, \bar{X} or \hat{X} , do you recommend using for hypothesis testing?

Answer: the power of test using \bar{X} is the probability of correctly rejecting H_0 : $\Pr(\bar{X} <$

$7.866|\mu = 8) = \Pr\left(\frac{\bar{X}-8}{\sqrt{\text{Var}(\bar{X})}} < \frac{7.866-8}{1.667}|\mu = 8) = \Pr(Z < -0.0803) \approx 0.5 - 0.0319 = 0.468$. **The power of test using \hat{X} is the probability of correctly rejecting H_0 : $\Pr(\bar{X} < 6.055|\mu = 8) = \Pr\left(\frac{\bar{X}-8}{\sqrt{\text{Var}(\bar{X})}} < \frac{6.055-8}{3.082}|\mu = 8) = \Pr(Z < -0.631) \approx 0.5 - 0.2357 = 0.264$. **Test using \bar{X} has a higher power than \hat{X} and, therefore, we recommend using \bar{X} for hypothesis testing. Grading Note: 2 points for each of (i) and (ii). 2 point for recommending \bar{X} with proper reasoning (If there is no reasoning, no point).****

7. (6 points) Suppose X_i for $i = 1, \dots, n$ is randomly drawn from normal distribution with mean μ and variance σ^2 . Let $\bar{X} = (1/n) \sum_{i=1}^n X_i$ be the sample average. Prove that $\text{Var}(\bar{X}) = \frac{\sigma^2}{n}$.

Answer: $\text{Var}(\bar{X}) = \text{Var}((1/n)(X_1+X_2+\dots+X_n)) = (1/n)^2 \text{Var}(X_1+X_2+\dots+X_n) = (1/n)^2 (\text{Var}(X_1) + \text{Var}(X_2) + \dots + \text{Var}(X_n)) = (1/n)^2 (n \times \sigma^2) = \frac{\sigma^2}{n}$, **where the third equality follows because $\text{Cov}(X_i, X_j) = 0$ if $i \neq j$ by random sampling.**

8. (6 points) Let X be a random variable with $E[X] = \mu$ and $\text{Var}(X) = \sigma^2$. Define $Z = \frac{X-\mu}{\sigma}$. Prove that $E[Z] = 0$ and $\text{Var}(Z) = 1$.

Answer: $E[Z] = E\left(\frac{X-\mu}{\sigma}\right) = \frac{1}{\sigma} E[X - \mu] = \frac{1}{\sigma} (E[X] - \mu) = \frac{1}{\sigma} \times 0 = 0$. $\text{Var}(Z) = \text{Var}\left(\frac{X-\mu}{\sigma}\right) = \frac{1}{\sigma^2} \text{Var}(X - \mu) = \frac{1}{\sigma^2} E[((X - \mu) - E(X - \mu))^2] = \frac{1}{\sigma^2} E[(X - \mu)^2] = \frac{1}{\sigma^2} \text{Var}(X) = 1$.

9. (6 points) The chi-square distribution with the r degree of freedom, denoted by $\chi^2(r)$, is characterized by the sum of r independent standard normally distributed random variables $Z_1^2, Z_2^2, \dots, Z_r^2$, where $Z_i \sim N(0, 1)$ and Z_i and Z_j are independent if $i \neq j$ for $i, j = 1, \dots, r$. Namely, $W = Z_1^2 + Z_2^2 + \dots + Z_r^2$ has a distribution that is $\chi^2(r)$. Prove that, if W has a $\chi^2(r)$ distribution, then $E[W] = r$.

Answer: $E[W] = E[Z_1^2 + Z_2^2 + \dots + Z_r^2] = E[Z_1^2] + E[Z_2^2] + \dots + E[Z_r^2] = 1 + 1 + \dots + 1 = r$, **where the third equality follows from $E[Z_i^2] = \text{Var}(Z_i) = 1$ because Z_i is a standard normal random variable with mean 0 and variance 1.**