

Identification and Estimation of Production Function with Unobserved Heterogeneity

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Motivation

- ▶ Estimation of production function and TFP is important in empirical applications (Empirical IO, Trade, Macro).
- ▶ In most empirical applications, the coefficients of Cobb-Douglas production function are assumed to be common across firms

$$y_{it} = \beta_0 + \beta_m m_{it} + \beta_\ell \ell_{it} + \beta_k k_{it} + \omega_{it} + \epsilon_{it}$$

→ $(\beta_m, \beta_\ell, \beta_k)$ is assumed to be common across i 's within “narrowly defined” industry

Is $(\beta_m, \beta_\ell, \beta_k)$ really common across firms?

Implications of Cobb-Douglas Production Function:

$$\frac{P_{M,t} M_{it}}{P_{Y,t} Y_{it}} \approx \beta_m$$

$$\frac{P_{M,t} M_{it}}{P_{M,t} M_{it} + W_t L_{it}} \approx \frac{\beta_m}{\beta_m + \beta_\ell}$$

$\left(\frac{P_{M,t} M_{it}}{P_{Y,t} Y_{it}} \right)_i$ over 25 yrs in Japanese Concrete Product and Electric Audio Industries, 30+ Workers

Figure: Concrete Product (2223)

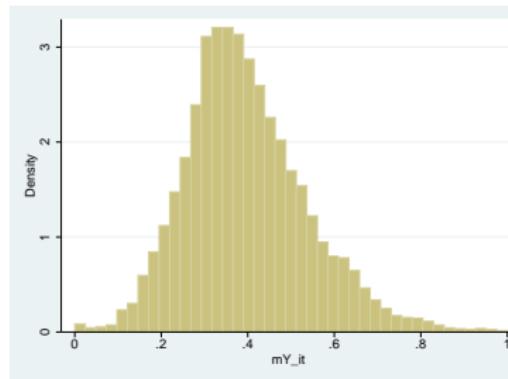
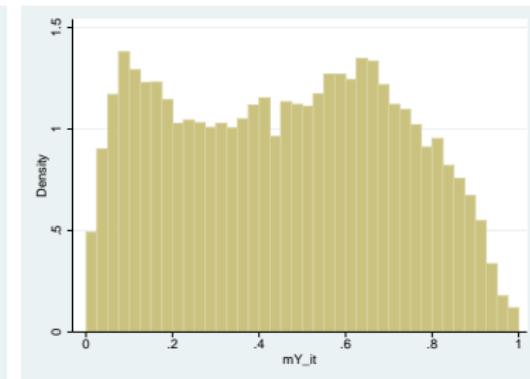


Figure: Electric Audio (2814)



The share of materials is **heterogenous** across firms and **persistent** over time within firm.

$$\left(\frac{P_{M,t} M_{it}}{P_{M,t} M_{it} + W_t L_{it}} \right)_i \text{ over 25 years}$$

Figure: Concrete Product (2223)

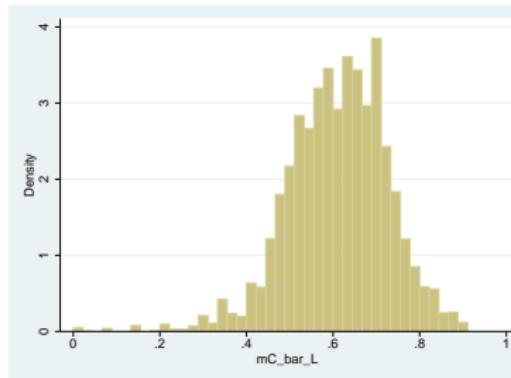
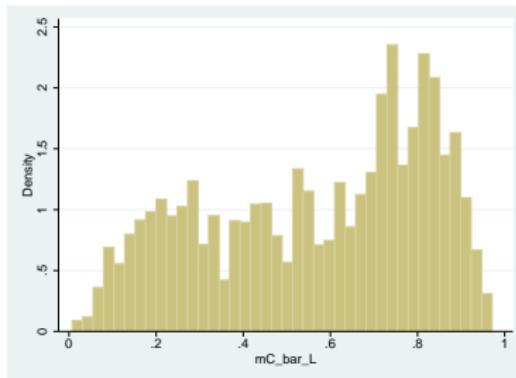


Figure: Electric Audio (2814)



⇒ The difference in markups is not the main reason for the heterogeneity in $\frac{PM_{it}}{PY_{it}}$.

$\left(\frac{P_{M,t} M_{it}}{P_{Y,t} Y_{it}} \right)_i$: 2-digit vs. 3-digit vs. 4-digit

Figure: 2-digit
Electric Parts, Device,
Circuit (28)

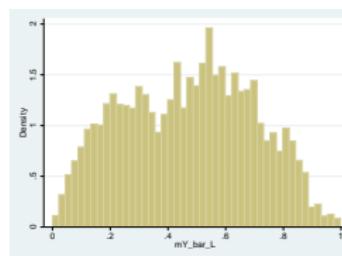


Figure: 3-digit
Electric Device (281)

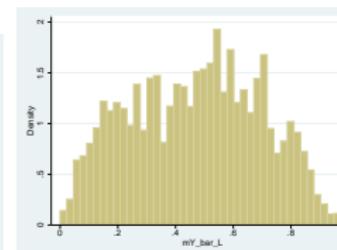
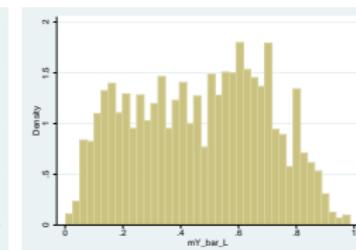


Figure: 4-digit
Electric Audio (2814)



Industry Code : Name	No. of Obs.	90-10 diff in $\left(\frac{PM_{it}}{PQ_{it}} \right)_i$	90-10 diff in $\left(\frac{PM_{it}}{PM_{it} + WL_{it}} \right)_i$
28: Electric Parts/Device/Circuit	33,467	0.62	0.66
281: Electric Device	21,712	0.62	0.67
2814: Electric Audio	12,441	0.63	0.67

The average difference between the 90th and the 10th percentiles at 2-, 3-, and 4-digit industry classifications

Industry Classifications	No. of Industries	Ave. 90-10 diff in $\left(\frac{PM_{it}}{PQ_{it}} \right)_i$	Ave. 90-10 diff in $\left(\frac{PM_{it}}{PM_{it} + WL_{it}} \right)_i$	Ave. No. of Obs.
2-digit	24	0.46	0.44	53,205
3-digit	149	0.42	0.41	8,570
4-digit	279	0.38	0.37	2,666

More general than Cobb-Douglas? (Electric Audio)

$$\frac{P_{M,t} M_{it}}{P_{Y,t} Y_{it}} = \text{2nd order polynomials of } (l_{it}, k_{it}, m_{it}) + e_{it}$$

$$\hat{\xi}_i := T^{-1} \sum_{t=1}^T \hat{e}_{it}$$

- ▶ 90-10 diff of $\hat{\xi}_i = 0.52$
- ▶ 90-10 diff of $\left(\frac{PM_{it}}{PQ_{it}} \right)_i = 0.63$
- ▶ Evidence for **persistent heterogeneity** under more general functions than Cobb-Douglas.

Issues

- ▶ Can we identify production function in the presence of unobserved heterogeneity?
- ▶ Can we estimate production function with random coefficients using a typical firm/plant-level panel data?

Issues in Estimation of Firm-Level Production Function

Simultaneity Bias

- ▶ Corr. between productivity & input factors (Marschak and Andrews, 1944)
- ▶ Control function/ “Proxy variable” approach: OP (Olley and Pakes, 1996), LP (Levinsohn and Petrin, 2003), Wooldridge (2009), ACF (Ackerberg, Caves and Frazer, 2015)
- ▶ Dynamic Panel: Arellano and Bond (1991), Blundell and Bond (1998)
- ▶ “Exogenous” Input Price as IV: Doraszelski and Jaumandreu (2018)

→ OP/LP is widely used in empirical analysis.

Identification of Firm-Level Production Function

Nonparametric Identification Problem

- ▶ Bond and Söderbom (2005) and ACF (2015): Colinearity of material and labor.
- ▶ GNR (Gandhi, Navarro, and Rivers, 2020): No exogenous variation to identify flexible input's elasticity after conditioning on quasi-fixed input factors.
- ▶ GNR exploit the first order condition w.r.t. flexible inputs for identification.

Production Function with Random Coefficients

- ▶ Mairesse and Griliches (1990), Van Biesebroeck (2003), Doraszelski and Jaumandreu (2018)
- ▶ Identification without input price instruments?
→ Unresolved issue.

Our Paper's contribution

- ▶ Non-parametric identification of production function with unobserved heterogeneity
- ▶ Estimation of production function with random coefficients
- ▶ Empirical application: Japanese Census of Manufacture

Nonparametric Identification

Setup

- ▶ Panel Data: $\{\{Y_{it}, \tilde{L}_{it}, K_{it}, M_{it}, B_{it}\}_{t=1}^T\}_{i=1}^n$
 - ▶ Y_{it} Output; \tilde{L}_{it} # of workers; K_{it} Capital stock; M_{it} Intermediate input; B_{it} Total wage bills
- ▶ K_{it} : Predetermined
- ▶ \tilde{L}_{it}, M_{it} : Flexibly chosen
- ▶ \mathcal{I}_{it} : Information available for choosing L_{it} & M_{it}

Production function and Capital

- ▶ J unobserved types with $\pi^j := \Pr(\text{type}=j)$

$$Y_{it} = e^{\omega_{it} + \epsilon_{it}} F_t^j(L_{it}, K_{it}, M_{it}), \quad \epsilon_{it} | \mathcal{I}_{it} \sim_{iid} g_{\epsilon,t}^j(\epsilon_{it})$$

$$\omega_{it} = h^j(\omega_{it-1}) + \eta_{it}, \quad \eta_{it} | \mathcal{I}_{it-1} \sim_{iid} g_\eta^j(\eta_{it})$$

- ▶ K_{it} predetermined with

$$P_t(K_{it} | \mathcal{I}_{t-1}, D_i = j) = P_t^j(K_{it} | K_{it-1}, \omega_{it-1})$$

Labor Input and Wage Bills

- L_{it} not directly observed but related to # of workers \tilde{L}_{it} as

$$L_{it} = e^{\psi_t^j} \tilde{L}_{it}$$

- Total Wage Bills:

$$B_{it} = e^{v_{it} + \zeta_{it}} P_{L,t} L_{it}, \quad v_{it} | \mathcal{I}_{it-1} \sim_{iid} g_v^j(v_{it})$$

Average wage B_{it}/\tilde{L}_{it}

⇒ Identification of unobserved quality/hours ψ_t^j

L and M are flexible inputs

- L_{it} & M_{it} are chosen flexibly before observing ϵ_{it} & ζ_{it} :

$$\begin{aligned}(M_{it}, L_{it}) &= (\mathbb{M}_t^j(K_{it}, \omega_{it}, v_{it}), \mathbb{L}_t^j(K_{it}, \omega_{it}, v_{it})) \\ &:= \operatorname{argmax}_{(M, L) \in \mathcal{M} \times \mathcal{L}} P_{Y,t} E^j[e^{\epsilon_{it}} | \mathcal{I}_{it}] e^{\omega_{it}} F_t^j(K_{it}, L, M) - P_{M,t} M \\ &\quad - E^j[e^{\zeta_{it}} | \mathcal{I}_{it}] e^{v_{it}} P_{L,t} L,\end{aligned}$$

- $(\mathbb{M}_t^j(K_{it}, \omega_{it}, v_{it}), \mathbb{L}_t^j(K_{it}, \omega_{it}, v_{it}))$ is invertible w.r.t. (ω_{it}, v_{it}) with probability one
- Correlation between (L_{it}, M_{it}) and $\omega_{it} \Rightarrow$ Endogeneity

First Order Conditions as Identifying Restrictions

- ▶ Material & Labor Share Equations:

$$\frac{P_{M,t} M_{it}}{\underbrace{P_{Y,t} Y_{it}}_{:=S_{it}^m}} = \frac{\partial F_t^j(L_{it}, K_{it}, M_{it}) / \partial M}{\underbrace{F_t^j(L_{it}, K_{it}, M_{it}) / M_{it}}_{:=G_{M,t}^j(L_{it}, K_{it}, M_{it})}} E[e^\epsilon | \text{type}=j] e^{-\epsilon_{it}}$$

$$\frac{B_{it}}{\underbrace{P_{Y,t} Y_{it}}_{:=S_{it}^\ell}} = \frac{\partial F_t^j(L_{it}, K_{it}, M_{it}) / \partial L}{\underbrace{F_t^j(L_{it}, K_{it}, M_{it}) / L_{it}}_{:=G_{L,t}^j(L_{it}, K_{it}, M_{it})}} \frac{E[e^\epsilon | \text{type}=j]}{E[e^\zeta | \text{type}=j]} e^{\zeta_{it} - \epsilon_{it}}$$

- ▶ Material/labor share equation ⇒ Source of Identification

Main Nonparametric Identification Results

Under regularity conditions, the model structure

$$\theta = \{g_{\epsilon\zeta,t}^j(\cdot), g_v^j(\cdot), g_\eta^j(\cdot), h^j(\cdot), \pi^j, \{G_{M,t}^j(\cdot), G_{L,t}^j(\cdot), F_t^j(\cdot)\}_{t=1}^T, P_{L,t}, \psi_t^j\}_{j=1}^J,$$

is nonparametrically identified from the joint distribution of
 $T = 4$ periods of the panel data,

$$\Pr(\{Y_t, B_t, M_t, \tilde{L}_t, K_t\}_{t=1}^4)$$

Or equivalently, consider the following data

$$\Pr(\{S_t^m, S_t^\ell, M_t, \tilde{L}_t, K_t\}_{t=1}^4).$$

Cobb-Douglas Case

$$\begin{aligned}y_t &= \beta_0^j + \beta_m^j m_t + \beta_\ell^j (\psi_t^j + \tilde{\ell}_t) + \beta_k^j k_t + \omega_t + \epsilon_t \\s_t^m &= \ln \beta_m^j + \ln E[e^\epsilon | \text{type}=j] - \epsilon_t \\s_t^\ell - s_t^m &= \ln(\beta_\ell^j / \beta_m^j) - \ln(E_t^j[e^\zeta]) + \zeta_t \\m_t - \tilde{\ell}_t &= \underbrace{\ln(P_{L,t}/P_{M,t})}_{:=\alpha_t} + \ln\left(\beta_m^j / \beta_\ell^j\right) + \ln\left(E_t^j[e^\zeta]\right) + \psi_t^j + v_t\end{aligned}$$

We may identify:

- ▶ $\{\beta_0^j, \beta_m^j, \beta_\ell^j, \beta_k^j, \psi_t^j, g_{\epsilon\zeta,t}^j(\cdot)\}_{j,t}$
- ▶ $\{\pi^j, h^j(\cdot), g_\eta^j(\cdot), g_v^j(\cdot)\}_j, \{\alpha_t\}_{t=1}^T$

from the distribution of $\{S_t, \tilde{X}_t\}_{t=1}^T$ for $T \geq 4$.

Estimation

Two-stage Parametric MLE: Additional Assumptions

- ▶ J unobserved types
- ▶ Cobb-Douglas Production Function
- ▶ Assume normal distribution:

$$(\epsilon_{it}, \zeta_{it}, v_{it})' \stackrel{d}{\sim} N(0, \Sigma^j)$$

$$\omega_{it} = \rho_\omega^j \omega_{it-1} + \eta_{it}, \quad \eta_i \stackrel{d}{\sim} N(0, \sigma_\eta^j)$$

$$k_{it}|(k_{it-1}, \omega_{it-1}) \stackrel{d}{\sim} N(\rho_{k0}^j + \rho_{kk}^j k_{it-1} + \rho_{k\omega}^j \omega_{it-1}, (\sigma_k^j)^2)$$

$$(k_{i1}, \omega_{i1}) \stackrel{d}{\sim} N(\mu_1^j, \Sigma_1^j).$$

Two-stage Parametric MLE

Probability Density Function of $(s_{it}, \tilde{x}_{it})_{t=1}^T$ for type j

$$f_t^j(\{s_{it}, \tilde{x}_{it}\}_{t=1}^T) = \underbrace{\prod_{t=1}^T f_t^j(s_{it}, \tilde{\ell}_{it} - m_{it}; \theta_1^j, \alpha_t)}_{=L_{1i}(\theta_1^j, \alpha)}$$

$$\times \underbrace{f_1^j(\tilde{x}_{i1} | \tilde{\ell}_{i1} - m_{i1}; \theta^j) \prod_{t=2}^T f_t^j(\tilde{x}_{it} | \tilde{\ell}_{it} - m_{it}, \tilde{x}_{it-1}; \theta^j)}_{=L_{2i}(\theta_1^j, \theta_2^j)},$$

Two-stage Parametric MLE

- ▶ First Stage:

$$\hat{\theta}_1 = \underset{(\pi^j, \theta_1^j)_j, \alpha}{\operatorname{argmax}} \sum_{i=1}^N \log \left(\sum_{j=1}^J \pi^j L_{1i}(\theta_1^j, \alpha) \right)$$

- ▶ Second Stage: Given $((\hat{\theta}_1^j)_j, \hat{\alpha})$,

$$\hat{\theta}_2 = \underset{\pi, (\theta_2^j)_j}{\operatorname{argmax}} \sum_{i=1}^N \log \left(\sum_{j=1}^J \pi^j L_{1i}(\hat{\theta}_1^j, \hat{\alpha}) L_{2i}(\hat{\theta}_1^j, \theta_2^j) \right)$$

Two-Step Parameteric MLE

- ▶ Advantage: Short T ($T \geq 4$), based on identification argument, type is identified from all data
- ▶ Disadvantage: Parametric assumptions, Computationally difficult due to local maxima.
- ▶ EM algorithm

Empirical Application: Japanese Census of Manufacture, 1986-2010

Data set

- ▶ Unbalanced panel data, Plants with 30+ workers
- ▶ Variable Construction
 - ▶ K_{it} : Fixed asset less land. Perpetual inventory method.
 - ▶ L_{it} : Number of employees
 - ▶ M_{it} : Material + Energy + Subcontracting expenses for consigned production
 - ▶ Y_{it} : Sales
- ▶ Electric Audio Equipment: 907 plants \times 4~25 years

Specification

- ▶ Cobb-Douglas (3 Types):

$$Y_i = \beta_0^j + \beta_m^j m_{it} + \beta_k^j k_{it} + \beta_\ell^j \ell_{it} + \omega_{it} + \epsilon_{it}, \quad \text{for } j = 1, 2, 3$$

- ▶ Unobserved hours or labor quality (2 × 3 = 6 Types):

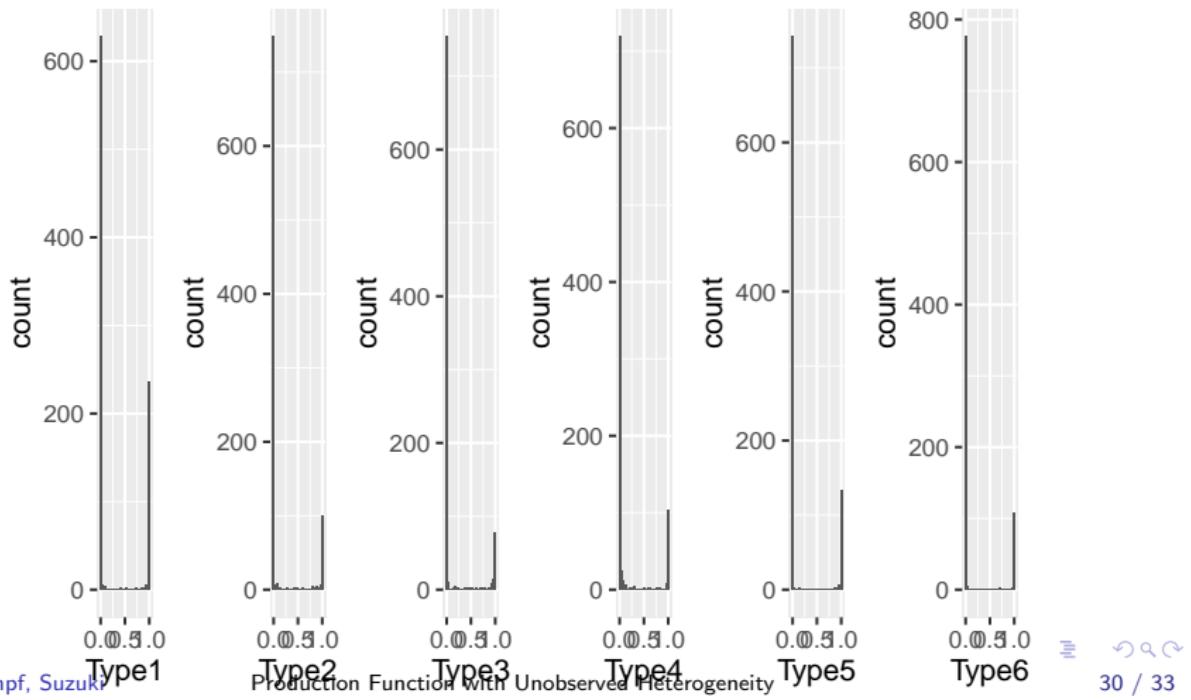
$$L_{it} = \psi^j \tilde{L}_{it} \quad \text{for } j = 1, 2, \dots, 6$$

Estimates of Production Function (Electric Audio)

	J = 1	J = 6					
		Type 1	Type 2	Type 3	Type 4	Type 5	Type 6
β_m^j	0.281 (0.008)	0.136 (0.007)		0.364 (0.022)		0.597 (0.010)	
β_ℓ^j	0.296 (0.006)	0.507 (0.012)		0.228 (0.012)		0.132 (0.006)	
β_k^j	0.076 (0.016)	0.122 (0.010)		0.185 (0.021)		0.202 (0.017)	
$\beta_m^j + \beta_\ell^j + \beta_k^j$	0.652	0.765		0.778		0.932	
β_k^j / β_ℓ^j	0.256	0.242		0.813		1.526	
ψ^j	0.000 (0.135)	-0.999 (0.139)	0.474	-0.258 (0.389)	0.201 (0.142)	-0.215 (0.115)	0.798 (—)
π	1.000 (0.023)	0.278 (0.017)	0.137	0.134 (0.163)	0.158 (0.149)	0.164 (0.015)	0.128 (0.014)

Posterior Probabilities for $J = 6$ Types (Electric Audio)

$$\hat{\pi}_i^j = \frac{\hat{\pi}^j L_i(\hat{\theta}^j)}{\sum_{k=1}^J \hat{\pi}^k L_i(\hat{\theta}^k)} \quad \text{for } j = 1, \dots, J.$$



The Effect of ω_{it} on Investment (Electric Audio)

$$I_{it}/K_{it} = \alpha_0 + \alpha_\omega^j \omega_{it} + \text{quadratic of } k_{it}$$

	$J = 1$	$J = 6$					
		Type 1	Type 2	Type 3	Type 4	Type 5	Type 6
α_ω^j	0.50 (0.11)	-0.19 (0.13)	-0.47 (1.26)	0.29 (0.07)	0.18 (0.34)	0.11 (0.02)	0.10 (0.03)
β_m^j	0.28	0.14		0.36		0.60	
β_k^j/β_ℓ^j	0.26	0.24		0.81		1.53	

The effect of ω_{it} on I_{it}/K_{it} is larger among firms with

- ▶ high material share
- ▶ high capital-labor ratio

Quantile Regression (Electric Audio)

$$I_{it}/K_{it} = \alpha_0 + \alpha_{\omega}^j \omega_{it} + \text{quadratic of } k_{it}$$

	$J = 1$	$J = 6$					
		Type 1	Type 2	Type 3	Type 4	Type 5	Type 6
$\alpha_{\omega}^j(0.10)$	0.00 (0.00)	0.00 (0.00)	0.00 (0.00)	0.01 (0.01)	0.00 (0.00)	-0.00 (0.01)	0.01 (0.01)
$\alpha_{\omega}^j(0.50)$	0.01 (0.00)	0.00 (0.00)	0.02 (0.00)	0.04 (0.01)	0.01 (0.00)	0.04 (0.01)	0.04 (0.01)
$\alpha_{\omega}^j(0.90)$	0.06 (0.00)	0.01 (0.01)	0.01 (0.02)	0.09 (0.03)	0.09 (0.02)	0.18 (0.05)	0.10 (0.04)
β_m^j	0.28	0.14		0.36		0.60	
β_k^j/β_{ℓ}^j	0.26	0.24		0.81		1.53	

Conclusions

- ▶ Evidence for heterogeneity beyond Hick's neutral term
- ▶ Non-parametric identification
- ▶ Maximum likelihood estimation
- ▶ Empirical applications to Japanese Census Data
- ▶ Heterogeneous investment/export elasticities across unobserved types